

talagrand's convex distance

The weighted Hamming distance is

$$d_\alpha(x, A) = \inf_{y \in A} d_\alpha(x, y) = \inf_{y \in A} \sum_{i: x_i \neq y_i} |\alpha_i|$$

where $\alpha = (\alpha_1, \dots, \alpha_n)$. The same argument as before gives

$$\mathbb{P} \left\{ d_\alpha(X, A) \geq t + \sqrt{\frac{\|\alpha\|^2}{2} \log \frac{1}{\mathbb{P}\{A\}}} \right\} \leq e^{-2t^2/\|\alpha\|^2},$$

This implies

$$\sup_{\alpha: \|\alpha\|=1} \min(\mathbb{P}\{A\}, \mathbb{P}\{d_\alpha(X, A) \geq t\}) \leq e^{-t^2/2}.$$

convex distance inequality

convex distance:

$$d_T(x, A) = \sup_{\alpha \in [0, \infty)^n : \|\alpha\|=1} d_\alpha(x, A) .$$

Talagrand's convex distance inequality:

$$\mathbb{P}\{A\} \mathbb{P}\{d_T(X, A) \geq t\} \leq e^{-t^2/4} .$$

Follows from the fact that $d_T(X, A)^2$ is $(4, 0)$ weakly self bounding (by a saddle point representation of d_T).

Talagrand's original proof was different.

convex lipschitz functions

For $\mathbf{A} \subset [0, 1]^n$ and $x \in [0, 1]^n$, define

$$D(x, \mathbf{A}) = \inf_{y \in \mathbf{A}} \|x - y\| .$$

If \mathbf{A} is convex, then

$$D(x, \mathbf{A}) \leq d_T(x, \mathbf{A}) .$$

Proof:

$$\begin{aligned} D(x, \mathbf{A}) &= \inf_{\nu \in \mathcal{M}(\mathbf{A})} \|x - \mathbb{E}_\nu Y\| \quad (\text{since } \mathbf{A} \text{ is convex}) \\ &\leq \inf_{\nu \in \mathcal{M}(\mathbf{A})} \sqrt{\sum_{j=1}^n (\mathbb{E}_\nu \mathbb{1}_{x_j \neq Y_j})^2} \quad (\text{since } x_j, Y_j \in [0, 1]) \\ &= \inf_{\nu \in \mathcal{M}(\mathbf{A})} \sup_{\alpha: \|\alpha\| \leq 1} \sum_{j=1}^n \alpha_j \mathbb{E}_\nu \mathbb{1}_{x_j \neq Y_j} \quad (\text{by Cauchy-Schwarz}) \\ &= d_T(x, \mathbf{A}) \quad (\text{by minimax theorem}) . \end{aligned}$$



John von Neumann (1903–1957)

convex lipschitz functions

Let $\mathbf{X} = (\mathbf{X}_1, \dots, \mathbf{X}_n)$ have independent components taking values in $[0, 1]$. Let $f : [0, 1]^n \rightarrow \mathbb{R}$ be quasi-convex such that $|f(\mathbf{x}) - f(\mathbf{y})| \leq \|\mathbf{x} - \mathbf{y}\|$. Then

$$\mathbb{P}\{f(\mathbf{X}) > \mathbb{M}f(\mathbf{X}) + t\} \leq 2e^{-t^2/4}$$

and

$$\mathbb{P}\{f(\mathbf{X}) < \mathbb{M}f(\mathbf{X}) - t\} \leq 2e^{-t^2/4}.$$

Proof: Let $A_s = \{\mathbf{x} : f(\mathbf{x}) \leq s\} \subset [0, 1]^n$. A_s is convex. Since f is Lipschitz,

$$f(\mathbf{x}) \leq s + D(\mathbf{x}, A_s) \leq s + d_T(\mathbf{x}, A_s),$$

By the convex distance inequality,

$$\mathbb{P}\{f(\mathbf{X}) \geq s + t\} \mathbb{P}\{f(\mathbf{X}) \leq s\} \leq e^{-t^2/4}.$$

Take $s = \mathbb{M}f(\mathbf{X})$ for the upper tail and $s = \mathbb{M}f(\mathbf{X}) - t$ for the lower tail.

empirical processes

Let \mathcal{T} be a countable index set.

For $i = 1, \dots, n$, let $\mathbf{X}_i = (\mathbf{X}_{i,s})_{s \in \mathcal{T}}$ be vectors of real-valued random variables. Assume that $\mathbf{X}_1, \dots, \mathbf{X}_n$ are independent.

The empirical process is $\sum_{i=1}^n \mathbf{X}_{i,s}, s \in \mathcal{T}$.

We study concentration of the supremum:

$$Z = \sup_{s \in \mathcal{T}} \sum_{i=1}^n \mathbf{X}_{i,s} .$$

empirical processes—the variance

We may use Efron-Stein: let

$$Z_i = \sup_{s \in \mathcal{T}} \sum_{j:j \neq i} X_{j,s}$$

and $\hat{s} \in \mathcal{T}$ be such that $Z = \sum_{i=1}^n X_{i,\hat{s}}$. Then

$$(Z - Z_i)_+ \leq (X_{i,\hat{s}})_+ \leq \sup_{s \in \mathcal{T}} |X_{i,s}|$$

so

$$\text{Var}(Z) \leq \mathbb{E} \sum_{i=1}^n (Z - Z_i)^2 \leq \mathbb{E} \sum_{i=1}^n \sup_{s \in \mathcal{T}} X_{i,s}^2.$$

empirical processes—the variance

A more clever use of Efron-Stein: suppose $\mathbb{E} \mathbf{X}_{i,s} = 0$.

Let $Z'_i = \sup_{s \in \mathcal{T}} \left(\sum_{j \neq i} \mathbf{X}_{j,s} + \mathbf{X}'_{i,s} \right)$. Note that

$$(Z - Z'_i)_+^2 \leq (\mathbf{X}_{i,\hat{s}} - \mathbf{X}'_{i,\hat{s}})^2.$$

By Efron-Stein,

$$\begin{aligned} \text{Var}(Z) &\leq \mathbb{E} \sum_{i=1}^n (Z - Z'_i)_+^2 \\ &\leq \mathbb{E} \sum_{i=1}^n \mathbb{E}' \left[(\mathbf{X}_{i,\hat{s}} - \mathbf{X}'_{i,\hat{s}})^2 \right] \\ &\leq \mathbb{E} \sum_{i=1}^n \left(\mathbf{X}_{i,\hat{s}}^2 + \mathbb{E}' [\mathbf{X}'_{i,\hat{s}}^2] \right) \\ &\leq \mathbb{E} \sup_{s \in \mathcal{T}} \sum_{i=1}^n \mathbf{X}_{i,s}^2 + \sup_{s \in \mathcal{T}} \sum_{i=1}^n \mathbb{E} \mathbf{X}_{i,s}^2. \end{aligned}$$

weak and strong variance

We have proved that

$$\text{Var}(\mathbf{Z}) \leq \mathbf{V} \quad \text{and} \quad \text{Var}(\mathbf{Z}) \leq \Sigma^2 + \sigma^2$$

where

$$\mathbf{V} = \sum_{i=1}^n \mathbb{E} \sup_{s \in \mathcal{T}} X_{i,s}^2 \quad \text{strong variance}$$

$$\Sigma^2 = \mathbb{E} \sup_{s \in \mathcal{T}} \sum_{i=1}^n X_{i,s}^2 \quad \text{weak variance}$$

$$\sigma^2 = \sup_{s \in \mathcal{T}} \sum_{i=1}^n \mathbb{E} X_{i,s}^2 \quad \text{wimpy variance}$$

$$\sigma^2 \leq \Sigma^2 \leq \mathbf{V} .$$

weak and strong variance

If $\mathbb{E}X_{i,s} = 0$ and $|X_{i,s}| \leq 1$, we also have, by symmetrization and contraction arguments,

$$\Sigma^2 \leq 8\mathbb{E}Z + \sigma^2$$

and therefore

$$\text{Var}(Z) \leq 8\mathbb{E}Z + 2\sigma^2.$$

If the X_i are also identically distributed, then

$$\text{Var}(Z) \leq 2\mathbb{E}Z + \sigma^2.$$

empirical processes—exponential inequalities

A Bernstein type inequality. “Talagrand’s inequality”.

Assume $\mathbb{E} \mathbf{X}_{i,s} = 0$, and $|\mathbf{X}_{i,s}| \leq 1$. For $t \geq 0$,

$$\mathbb{P}\{\mathbf{Z} \geq \mathbb{E} \mathbf{Z} + t\} \leq \exp\left(-\frac{t^2}{2(2(\Sigma^2 + \sigma^2) + t)}\right).$$

proof.

For each $i = 1, \dots, n$, let $Z'_i = \sup_{s \in \mathcal{T}} (X'_{i,s} + \sum_{j \neq i} X_{j,s})$. We already proved that

$$\sum_{i=1}^n \mathbb{E}'(Z - Z'_i)_+^2 \leq \sup_{s \in \mathcal{T}} \sum_{i=1}^n X_{i,s}^2 + \sigma^2 \stackrel{\text{def.}}{=} W + \sigma^2 .$$

By the exponential Efron-Stein inequality, for $\lambda \in [0, 1]$,

$$\log \mathbb{E} e^{\lambda(Z - \mathbb{E} Z)} \leq \frac{\lambda}{1 - \lambda} \log \mathbb{E} e^{\lambda(W + \sigma^2)} .$$

W is a self-bounding function, so

$$\log \mathbb{E} e^{\lambda W} \leq \Sigma^2 (e^\lambda - 1) .$$

Putting things together implies the inequality.

bousquet's inequality

A Bennett type inequality with the right constant.

Assume $\mathbf{X}_1, \dots, \mathbf{X}_n$ are i.i.d. with $\mathbb{E}\mathbf{X}_{i,s} = 0$ and $\mathbf{X}_{i,s} \leq 1$.
For all $t \geq 0$,

$$\mathbb{P}\{\mathbf{Z} \geq \mathbb{E}\mathbf{Z} + t\} \leq e^{-\nu h(t/\nu)} .$$

where $\nu = 2\mathbb{E}\mathbf{Z} + \sigma^2$ and $h(u) = (1+u)\log(1+u) - u$.
In particular,

$$\mathbb{P}\{\mathbf{Z} \geq \mathbb{E}\mathbf{Z} + t\} \leq \exp\left(-\frac{t^2}{2(\nu + t/3)}\right) .$$

ϕ entropies

For a convex function ϕ on $[0, \infty)$, the ϕ -entropy of $Z \geq 0$ is

$$H_\phi(Z) = \mathbb{E}[\phi(Z)] - \phi(\mathbb{E}[Z]).$$

H_ϕ is subadditive:

$$H_\phi(Z) \leq \sum_{i=1}^n \mathbb{E} \left[\mathbb{E} \left[\phi(Z) | X^{(i)} \right] - \phi \left(\mathbb{E} [Z | X^{(i)}] \right) \right]$$

if (and only if) ϕ is twice differentiable on $(0, \infty)$, and either ϕ is affine strictly positive and $1/\phi''$ is concave.

$\phi(x) = x^2$ corresponds to Efron-Stein.

$x \log x$ is subadditivity of entropy.

We may consider $\phi(x) = x^p$ for $p \in (1, 2]$.

generalized efron-stein

Define

$$Z'_i = f(X_1, \dots, X_{i-1}, X'_i, X_{i+1}, \dots, X_n) ,$$

$$V^+ = \sum_{i=1}^n (Z - Z'_i)_+^2 .$$

For $q \geq 2$ and $q/2 \leq \alpha \leq q - 1$,

$$\begin{aligned} & \mathbb{E} [(Z - \mathbb{E}Z)_+^q] \\ & \leq \mathbb{E} [(Z - \mathbb{E}Z)_+^\alpha]^{\alpha/q} + \alpha(\alpha - q) \mathbb{E} [V^+ (Z - \mathbb{E}Z)_+^{q-2}] \end{aligned}$$

and similarly for $\mathbb{E} [(Z - \mathbb{E}Z)_-^q]$.

moment inequalities

We may solve the recursions, for $\mathbf{q} \geq 2$.

If $\mathbf{V}^+ \leq \mathbf{c}$ for some constant $\mathbf{c} \geq 0$, then for all integers $\mathbf{q} \geq 2$,

$$(\mathbb{E} [(\mathbf{Z} - \mathbb{E}\mathbf{Z})_+^\mathbf{q}])^{1/\mathbf{q}} \leq \sqrt{\mathbf{Kqc}} ,$$

where $\mathbf{K} = 1 / (\mathbf{e} - \sqrt{\mathbf{e}}) < 0.935$.

More generally,

$$(\mathbb{E} [(\mathbf{Z} - \mathbb{E}\mathbf{Z})_+^\mathbf{q}])^{1/\mathbf{q}} \leq 1.6\sqrt{\mathbf{q}} \left(\mathbb{E} [\mathbf{V}^{+\mathbf{q}/2}] \right)^{1/\mathbf{q}} .$$

sums: khinchine's inequality

Let $\mathbf{X}_1, \dots, \mathbf{X}_n$ be independent Rademacher variables and $Z = \sum_{i=1}^n a_i \mathbf{X}_i$. For any integer $q \geq 2$,

$$(\mathbb{E}[Z_+^q])^{1/q} \leq \sqrt{2Kq} \sqrt{\sum_{i=1}^n a_i^2}$$

Proof:

$$V^+ = \sum_{i=1}^n \mathbb{E}[(a_i(\mathbf{X}_i - \mathbf{X}'_i))_+^2 \mid \mathbf{X}_i] = 2 \sum_{i=1}^n a_i^2 \mathbb{1}_{a_i \mathbf{X}_i > 0} \leq 2 \sum_{i=1}^n a_i^2 ,$$



Aleksandr Khinchin
(1894–1959)

sums: rosenthal's inequality

Let $\mathbf{X}_1, \dots, \mathbf{X}_n$ be independent real-valued random variables with $\mathbb{E}\mathbf{X}_i = 0$. Define

$$Z = \sum_{i=1}^n \mathbf{X}_i, \quad \sigma^2 = \sum_{i=1}^n \mathbb{E}\mathbf{X}_i^2, \quad Y = \max_{i=1, \dots, n} |\mathbf{X}_i|.$$

Then for any integer $q \geq 2$,

$$(\mathbb{E}[Z_+^q])^{1/q} \leq \sigma \sqrt{10q} + 3q (\mathbb{E}[Y_+^q])^{1/q} .$$

influences

If $\mathbf{A} \subset \{-1, 1\}^n$ and $\mathbf{X} = (\mathbf{X}_1, \dots, \mathbf{X}_n)$ is uniform, the influence of the i -th variable is

$$I_i(\mathbf{A}) = \mathbb{P} \{ \mathbb{1}_{\mathbf{X} \in \mathbf{A}} \neq \mathbb{1}_{\mathbf{X}^{(i)} \in \mathbf{A}} \}$$

where $\mathbf{X}^{(i)} = (\mathbf{X}_1, \dots, \mathbf{X}_{i-1}, 1 - \mathbf{X}_i, \mathbf{X}_{i+1}, \dots, \mathbf{X}_n)$.

The total influence is

$$I(\mathbf{A}) = \sum_{i=1}^n I_i(\mathbf{A}) .$$

Note that

$$I(\mathbf{A}) = 2^{-(n-1)} |\partial_E(\mathbf{A})| .$$

influences: examples

dictatorship: $\mathbf{A} = \{\mathbf{x} : x_1 = 1\}$. $I(\mathbf{A}) = 1$.

parity: $\mathbf{A} = \{\mathbf{x} : \sum_i \mathbb{1}_{x_i=1} \text{ is even}\}$. $I(\mathbf{A}) = n$.

majority: $\mathbf{A} = \{\mathbf{x} : \sum_i x_i > 0\}$. $I(\mathbf{A}) \approx \sqrt{2n/\pi}$.

by Efron-Stein, $P(\mathbf{A})(1 - P(\mathbf{A})) \leq \frac{I(\mathbf{A})}{4}$

so dictatorship has smallest total influence (if $P(\mathbf{A}) = 1/2$).

improved efron-stein on the hypercube

Recall that for any $\mathbf{f} : \{-1, 1\}^n \rightarrow \mathbb{R}$ under the uniform distribution,

$$\text{Ent}(\mathbf{f}^2) \leq 2\mathcal{E}(\mathbf{f})$$

where $\text{Ent}(\mathbf{f}^2) = \mathbf{E} [\mathbf{f}^2 \log(\mathbf{f}^2)] - \mathbf{E} [\mathbf{f}^2] \log \mathbf{E} [\mathbf{f}^2]$ and

$$\mathcal{E}(\mathbf{f}) = \frac{1}{4}\mathbb{E} \left[\sum_{i=1}^n \left(\mathbf{f}(\mathbf{X}) - \mathbf{f}(\bar{\mathbf{X}}^{(i)}) \right)^2 \right]$$

This implies, for any non-negative $\mathbf{f} : \{-1, 1\}^n \rightarrow [0, \infty)$,

$$\mathbf{E} [\mathbf{f}^2] \log \frac{\mathbf{E} [\mathbf{f}^2]}{\mathbf{E} [\mathbf{f}]^2} \leq 2\mathcal{E}(\mathbf{f}) .$$

improved efron-stein on the hypercube

Recall the Doob-martingale representation

$$f(X) - Ef = \sum_{i=1}^n \Delta_i.$$

One easily sees that

$$\mathcal{E}(f) = \sum_{i=1}^n \mathcal{E}(\Delta_i).$$

But then, by the previous lemma,

$$\begin{aligned} \mathcal{E}(f) &\geq \sum_{j=1}^n \mathcal{E}(|\Delta_j|) \geq \frac{1}{2} \sum_{j=1}^n E[\Delta_j^2] \log \frac{E[\Delta_j^2]}{(E|\Delta_j|)^2} \\ &= -\frac{1}{2} \text{Var}(f) \sum_{j=1}^n \frac{E[\Delta_j^2]}{\text{Var}(f)} \log \frac{(E|\Delta_j|)^2}{E[\Delta_j^2]} \\ &\geq -\frac{1}{2} \text{Var}(f) \log \frac{\sum_{j=1}^n (E|\Delta_j|)^2}{\text{Var}(f)} \end{aligned}$$

improved efron-stein on the hypercube

We obtained that for any $\mathbf{f} : \{-1, 1\}^n \rightarrow \mathbb{R}$,

$$\text{Var}(\mathbf{f}) \log \frac{\text{Var}(\mathbf{f})}{\sum_{j=1}^n (\mathbf{E}|\Delta_j|)^2} \leq 2\mathcal{E}(\mathbf{f}) .$$

(Falik and Samorodnitsky, 2007; Rossignol, 2006; Talagrand (1994)).

“Slightly” better than Efron-Stein.

Use this for $\mathbf{f}(\mathbf{x}) = \mathbf{1}_{\mathbf{x} \in \mathbf{A}}$ for $\mathbf{A} \subset \{-1, 1\}^n$:

$$\mathbf{P}(\mathbf{A})(1 - \mathbf{P}(\mathbf{A})) \log \frac{4\mathbf{P}(\mathbf{A})(1 - \mathbf{P}(\mathbf{A}))}{\sum_i I_i(\mathbf{A})^2} \leq \frac{I(\mathbf{A})}{4}$$

kahn, kalai, linial

Corollary: (Kahn, Kalai, Linial, 1988).

$$\max_i I_i(\mathbf{A}) \geq \frac{P(\mathbf{A})(1 - P(\mathbf{A})) \log n}{n}$$

If the influences are equal,

$$I(\mathbf{A}) \geq P(\mathbf{A})(1 - P(\mathbf{A})) \log n$$

Another corollary: (Friedgut, 1998).

If $I(\mathbf{A}) \leq c$, \mathbf{A} (basically) depends on a bounded number of variables. \mathbf{A} is a “junta.”

threshold phenomena

Let $\mathbf{A} \subset \{-1, 1\}^n$ be a monotone set and let $\mathbf{X} = (\mathbf{X}_1, \dots, \mathbf{X}_n)$ be such that

$$\mathbb{P}\{\mathbf{X}_i = 1\} = p \quad \mathbb{P}\{\mathbf{X}_i = -1\} = 1 - p$$

$$P_p(\mathbf{A}) = \sum_{x \in \mathbf{A}} p^{\|x\|} (1 - p)^{n - \|x\|}$$

is an increasing function of $p \in [0, 1]$.

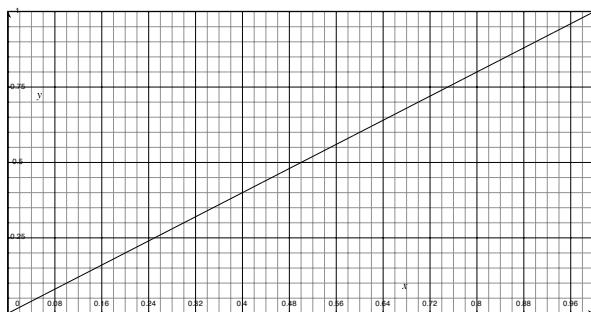
Let p_a be such that $P_{p_a}(\mathbf{A}) = a$.

Critical value = $p_{1/2}$

Threshold width: $p_{1-\varepsilon} - p_\varepsilon$

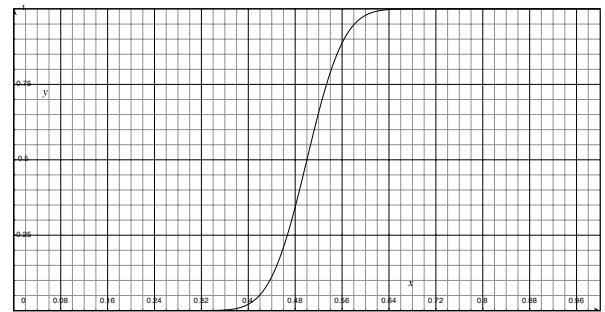
two (extreme) examples

dictatorship



threshold width = $1 - 2\epsilon$

majority (with $n = 101$)



$\leq \sqrt{\log(1/\epsilon)/(2n)}$

In what cases do we have a quick transition?

russo's lemma

If \mathbf{A} is monotone,

$$\frac{dP_p(\mathbf{A})}{dp} = I^{(p)}(\mathbf{A})$$

The Kahn, Kalai, Linial result, generalized for $p \neq 1/2$, implies that

if \mathbf{A} is such that $I_1^{(p)} = I_2^{(p)} = \dots = I_n^{(p)}$, then

$$p_{1-\varepsilon} - p_\varepsilon = O\left(\frac{\log \frac{1}{\varepsilon}}{\log n}\right)$$

On the other hand, if $p_{3/4} - p_{1/4} \geq c$ then \mathbf{A} is (basically) a junta.

books

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